

Acceptance and commitment therapy as guided self-help for psychological distress and positive mental health: a randomized controlled trial

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Background. In order to reduce the high prevalence of depression, early interventions for people at risk of depression are warranted. This study evaluated the effectiveness of an early guided self-help programme based on acceptance and commitment therapy (ACT) for reducing depressive symptomatology.

Method. Participants with mild to moderate depressive symptomatology were recruited from the general population and randomized to the self-help programme with extensive email support ($n=125$), the self-help programme with minimal email support ($n=125$) or to a waiting list control group ($n=126$). Participants completed measures before and after the intervention to assess depression, anxiety, fatigue, experiential avoidance, positive mental health and mindfulness. Participants in the experimental conditions also completed these measures at a 3-month follow-up.

Results. In the experimental conditions significant reductions in depression, anxiety, fatigue, experiential avoidance and improvements in positive mental health and mindfulness were found, compared with the waiting list condition (effect sizes Cohen's $d=0.51-1.00$). These effects were sustained at the 3-month follow-up. There were no significant differences between the experimental conditions on the outcome measures.

Conclusions. The ACT-based self-help programme with minimal email support is effective for people with mild to moderate depressive symptomatology.

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Key words: Acceptance and Commitment Therapy, depression, positive mental health, randomized controlled trial, self-help.

Introduction

In order to reduce the high prevalence of depression, early interventions for people at risk are warranted (Andrews *et al.* 2004; Cuijpers *et al.* 2008). Although such interventions are available and effective in reducing depressive symptoms and preventing the onset of depression, only a small proportion of people participate. Possible ways to increase the participation rate is to reduce the stigma associated with mental illnesses and to focus the intervention on depression as well as on other mental illnesses (Cuijpers *et al.* 2010). Early interventions should meet the following criteria: implementation within a positive framework to avoid the stigma associated with depression; targeting generic risk factors for multiple mental illnesses; and promoting positive mental health as well

as reducing psychological distress, because a sound mental health prevents the onset of clinical disorders (Keyes *et al.* 2010; Wood & Joseph, 2010). In this study an early intervention, based on acceptance and commitment therapy (ACT) and meeting these criteria, was offered to adults with mild to moderate depressive symptomatology as a guided self-help programme consisting of a manual 'Living to the full' (a positive frame) and email support by a counsellor. The intervention targets both experiential avoidance (EA) that can be considered a generic risk factor for mental illnesses (Biglan *et al.* 2008) and positive mental health. A group format of this intervention has shown promising effects (Fledderus *et al.* 2010; Bohlmeijer *et al.* 2011a). In this study two different levels of email support by a counsellor were explored (minimal and extensive) to assess optimal treatment intensity. We hypothesized that both ACT interventions would show superior effects in reducing depressive symptoms and other mental distress (anxiety, fatigue) and in increasing positive mental health when compared with a waiting list. Furthermore, we expected

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positive effects in the process measures EA and mindfulness.

Method

Participants

In September 2009, participants were recruited through advertisements in Dutch newspapers. Inclusion criteria were an age of 18 years or older and mild to moderate depressive symptoms [>10 and <39 on the Center of Epidemiological Studies – depression scale (CES-D); Radloff, 1977] and anxiety symptoms [>3 and <15 on the Hospital Anxiety and Depression Scale – anxiety subscale (HADS-A); Zigmond & Snaith, 1983]. People with severe depressive symptomatology and/or anxiety (more than 1 standard deviation above the population mean on the CES-D (cut-off score ≥ 39 ; Bouma *et al.* 1995) and/or HADS-A (cut-off score ≥ 15 ; Olsson *et al.* 2005) were excluded, because severe distress would require more intensive individual diagnostics and treatment. For the remaining participants it was checked who were still responding positively to a screener for a depressive disorder [Web Screening Questionnaire (WSQ; Donker *et al.* 2009) Q1 ≥ 6 and Q2 = 1]. As the WSQ yields a high number of false positives (Donker *et al.* 2009), those who were screened as having a depressive disorder underwent a telephone interview that employed the depressive episode module of the Mini International Neuropsychiatric Interview (MINI; Sheehan *et al.* 1998). People whom the MINI diagnosed as having a severe depressive episode were excluded.

Other exclusion criteria were: (a) few depressive symptoms (≤ 10 on the CES-D) and/or anxiety (≤ 3 on the HADS-A); (b) receiving psychological or psychopharmacological treatment within the last 3 months; and (c) high suicide risk (Q15 = 3 on the WSQ).

Procedure

A total of 625 people responded to the advertisements and received an information sheet explaining the study and an informed consent form. This was signed by 507 people who then received a screening questionnaire comprising the CES-D, HADS-A and WSQ. First, 54 respondents were excluded because they had severe depression and/or anxiety according to their scores on CES-D and HADS-A. They were advised to contact their general practitioner. Second, 44 respondents were diagnosed by the WSQ as having a depressive disorder and subsequently underwent a telephone interview using the MINI. These interviews were conducted by Masters students of psychology who were being trained and supervised by a clinical psychologist (K.M.G.S.). Of the 43 respondents

(one respondent could not be contacted), two were diagnosed with a severe depressive episode and were excluded and advised to contact their general practitioner. A further 75 respondents were excluded because they had few depression and/or anxiety symptoms ($n=58$), did not complete the screening questionnaire ($n=15$), could not be contacted for the interview ($n=1$) or currently received psychological treatment ($n=1$). Hence a total of 376 participants were randomly assigned to the following three conditions: the ACT intervention with minimal email support (ACT-M; $n=125$), the same intervention with extensive email support (ACT-E; $n=125$) or to a waiting list (W-L; $n=126$). A block randomization was carried out for the three groups with stratification on gender and age (<50 and ≥ 50 years), using computer-generated random sequences of numbers.

Table 1 shows an overview of the participants' characteristics. Their mean age was 42 years (range 18–73 years). The majority was female (70%) and of Dutch origin (93%). Most of the participants had a high level of education (86%), a paid job (76%) and were not married (47%).

Power analysis

A sample size of 100 participants per condition at post-intervention was needed to detect an effect size of 0.40 (Cohen's d) for the primary outcome with a statistical power of $(1 - \beta) = 0.80$ in a two-tailed test ($p < 0.05$). Taking into account a drop-out rate of 20%, we needed to recruit 375 people.

Measures

All participants completed measures on two occasions: at baseline (T0) and at post-intervention at 9 weeks (T1; directly after the intervention). Those assigned to the experimental conditions completed a third measure at the 3-month follow-up (T2; 5 months after baseline). The W-L group received the intervention after the waiting list period of 9 weeks. All questionnaires were administered online. The primary outcome measure was the level of depressive symptomatology as measured with the CES-D (20 items, score 0–60). Higher scores mean more depressive symptoms (Radloff, 1997; Haringsma *et al.* 2004).

Secondary outcome measures were anxiety, fatigue and a positive mental health. Anxiety was measured with the HADS-A (seven items, score 0–21) for assessing the presence and severity of anxious symptoms. Higher scores mean more anxiety symptoms (Spinoven *et al.* 1997; Zigmond & Snaith, 2003). The subjective fatigue subscale of the Checklist Individual Strength was used to assess the severity of fatigue (eight items, score 8–56). A high score indicates severe

Table 1. Characteristics of participants

Characteristic	ACT-E (<i>n</i> = 125)	ACT-M (<i>n</i> = 125)	W-L (<i>n</i> = 126)
Gender, <i>n</i>			
Female	87	87	88
Male	38	38	38
Mean age, years (s.d.)	42.64 (10.96)	42.35 (11.09)	42.47 (11.29)
Marital status, <i>n</i>			
Married	58	49	57
Divorced	8	14	10
Widowed	4	0	0
Unmarried	54	62	59
Race, <i>n</i>			
Dutch	114	119	116
Other	11	6	10
Education, <i>n</i>			
High	112	101	112
Middle	12	20	13
Low	1	4	1
Daily activities, <i>n</i>			
Paid job	91	94	100
No job	33	31	26

ACT-E, Acceptance and commitment therapy with extensive email support;
 ACT-M, acceptance and commitment therapy with minimal email support;
 W-L, waiting list control; s.d., standard deviation.

fatigue (Vercoulen *et al.* 1999). Positive mental health was assessed with the Mental Health Continuum short form that measures emotional well-being (three items), social well-being (five items) and psychological well-being (six items). Participants were asked to rate how often they had experienced feelings of well-being in the past month on a scale ranging from 1 (never) to 6 (daily). A mean score across the individual items was computed for each component of well-being. Higher scores indicate greater emotional, social and psychological well-being (Keyes, 2005; Lamers *et al.* 2010).

Process measures included assessments of EA and mindfulness. The Acceptance and Action Questionnaire-II was used to measure the ability to accept aversive internal experiences and to pursue values in the presence of these experiences (10 items, score 10–70). Higher scores indicated lower levels of EA (Jacobs *et al.* 2008; Bond *et al.*, 2011). The Five Facet Mindfulness Questionnaire (FFMQ) was used to measure five facets of mindfulness: observing (eight items), describing (eight items), acting with awareness (eight items), non-judging (eight items) and non-reactivity (seven items) (all scores range from 8–40, except non-reactivity facet range from 7 to 35). Higher scores indicate more mindfulness (Baer *et al.* 2006; Bohlmeijer *et al.* 2011*b*). All outcome measures showed good psychometric properties.

To evaluate the satisfaction of the participants after the intervention, the Client Satisfaction Questionnaire (CSQ-8) was used which consists of eight items on a scale from 1 (very negative) to 4 (very positive). A mean score across individual items was computed (Attkisson & Zwick, 1982; De Brey, 1983). Also a question was included on how the participants evaluated the programme on a scale from 1 to 10.

Counsellors

The email support was provided by five psychology Masters students of the University of Twente. They received a 2-day workshop from a clinical psychologist with ample experience in ACT. The student counsellors studied the self-help manual and performed all the exercises included in the manual. In the workshop they learned the processes of ACT and practised writing emails in the role of both client and counsellor. Each student provided 25 participants with minimal email support and 25 participants with extensive email support during a period of 9 weeks. The emails were supervised by a clinical psychologist. On average the counsellor spent 9 min on the extensive email support and 3 min on the minimal email support.

Intervention and email support

Participants of the experimental conditions received the self-help book 'Living to the full' (Bohlmeijer &

Hulsbergen, 2008) by regular mail. The book comprises nine modules, divided into three parts. The participants were instructed to complete one module per week. The modules are based on six core processes of ACT that together promote psychological flexibility: acceptance (active and aware embracement of aversive internal experiences), cognitive defusion (creating a context in which undesirable functions of thoughts disappear), contact with the present moment, self as context (experiencing that one is more than one's thoughts, feelings and experiences), choosing values in different life domains, and commitment to choices on the basis of these values (Hayes *et al.* 2006).

In the first part of the book, participants reflect on their avoidance and control strategies and whether these are effective in the long run. In the second part, participants learn how to come into contact with their present experiences without trying to avoid or control them. Cognitive defusion and experiencing self as context are practised. In the third part, the focus is on becoming aware of the most important personal values and making decisions based on these values. Each module uses experiential exercises and metaphors for illustrating the processes of ACT. Furthermore, the participants were asked to do daily mindfulness exercises, based on mindfulness-based stress reduction (Kabat-Zinn, 1990, 1994). The mindfulness exercises lasted on average 10–15 min and were on an audio compact disc which was provided with the book.

Both experimental conditions received the same frequency of email support. The participants received a standardized email from the counsellors every week on the same day concerning the module that they had carried out in the previous week. Participants were expected to respond within 2 days, after which they received a feedback email from the counsellor. In the minimal email support the standardized email comprised questions on their progress (e.g. did you perform all the exercises?), which was met with positive and encouraging feedback from the counsellor. In the extensive email support the standardized email also comprised questions on what they had discovered or experienced in the previous week (e.g. what did you experience when you performed the exercises?). Furthermore, participants were allowed to ask a question on the text or exercises. The counsellor responded both to the question and to the experiences of the participants and offered advice and instruction with regard to the text and exercises.

Participants assigned to the W-L group received the self-help book by regular mail after the post-intervention assessment. Email counselling was not provided.

Statistical analyses

The statistical analyses were done using PASW 18 (Predictive Analytics Software; IBM, USA). One-way analysis of variance (ANOVA) and χ^2 tests showed that there were no significant differences at baseline between the experimental conditions and the W-L group for any of the demographic variables or outcome measures, indicating successful randomization.

Intention-to-treat analyses were conducted with the use of PASW Missing Value Analysis (SPSS Inc., USA) to impute all missing data on the continuous measures with the expectation-maximization method. This method computes missing values based on maximum likelihood estimates using observed data in an iterative process (Dempster *et al.* 1977). The total percentage of missing data was 7.9% due to unanswered items (0.7%) or incomplete assessments (7.2%). These missing values at baseline, post-intervention and follow-up were imputed. A comparison of results based on the imputed intention-to-treat sample *versus* the observed data revealed similar outcomes. Therefore, only the results from the intention-to-treat analyses are reported. Comparisons were two-tailed and interpreted with a significance of $p < 0.050$.

To examine differences between the conditions on all the outcome measures, a 3 (group) \times 2 (time) ANOVA was used. In the case of significant time \times group interactions, the *post hoc* test Tukey's honestly significant difference (HSD) test was used. To examine the change from post-intervention to follow-up in the experimental conditions, paired *t* tests and a 2 (group) \times 2 (time) ANOVA were used. To identify whether gender, educational level (high *v.* low), the level of depression, anxiety and EA at baseline moderated the effect of the intervention on the change of depressive symptomatology from baseline to post-intervention, regression analyses were used. The medians on the scores on depression, anxiety and EA at baseline were used as cut-off scores for recoding them as dichotomous variables. In a linear regression model the intervention dummy variable (experimental conditions = 1 *v.* W-L = 0), interaction term with the potential moderator as a dichotomous variable and their main effects were entered.

Effect sizes at post-intervention were calculated with Cohen's *d* using the means and the pooled standard deviations of the measurements of the conditions (effect size of 0.56–1.2 was considered large, 0.33–0.55 as moderate, and less than 0.33 as small) (Lipsey *et al.* 1993).

With the Jacobson and Truax methodology, the proportion of participants was determined who made a clinically significant change on the CES-D from

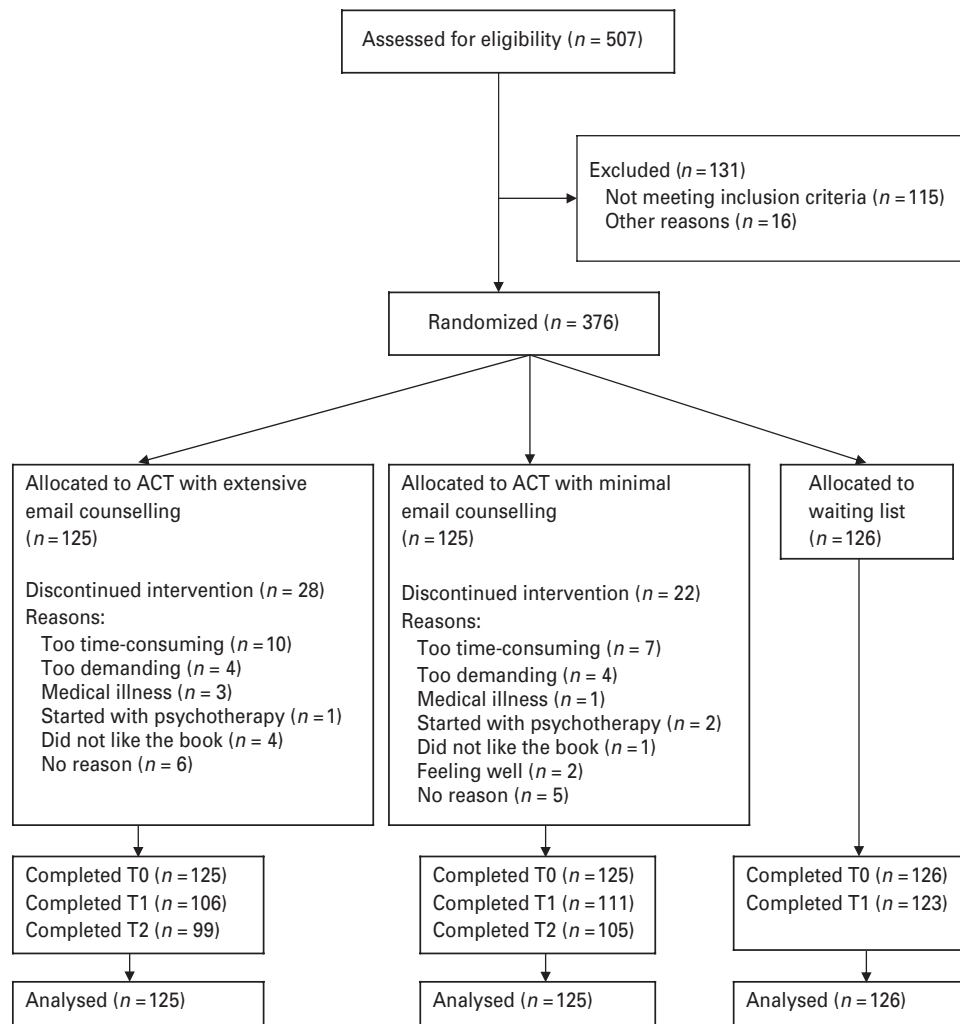


Fig. 1. Participant flow. ACT, Acceptance and commitment therapy; T0, baseline; T1, post-intervention at 9 weeks (directly after the intervention); T2, 3-month follow-up (5 months after baseline).

baseline to post-treatment (Jacobson & Traux, 1991). First, the reliable change was calculated with the reliable change index (a reduction of ≥ 12 points). Second, because we studied a population with mild to moderate depressive symptomatology and people with mild to moderate depressive disorders, the scores below one standard deviation of the mean pre-test scores on the CES-D were considered the recovery criterion (a score ≤ 16). The score of 16 also has been found in previous research as the cut-off score to indicate the presence of clinically relevant depressive symptoms (Beekman *et al.* 1997; Smit *et al.* 2006). A clinically significant change on the CES-D is thus defined as having a reliable change (a reduction of at least 12 points) between the measurements and thereby crossing the cut-off of 16. Participants that had a clinically relevant change were either coded 1 (implying favourable treatment response, 'success') or 0 ('failure'). The binary outcome was used to calculate

the odds ratio (OR) using logistic regression. Based on the clinically significant change proportions, the number needed to treat (NNT) was calculated (Cook & Sackett, 1995).

Results

Drop-out

In Fig. 1 the flow of participants is given. At post-intervention, data were available for 340 participants (drop-out rate 9.6%) and at follow-up for 204 participants (drop-out rate 18.4%). There were no significant differences at baseline on all the measures between participants who completed the assessments and those who did not.

Intervention adherence and participants' satisfaction

There were 22 participants (17.6%) in the ACT-M group and 28 (22.4%) in the ACT-E group that did not

fully adhere to the intervention. The main reasons given for non-completion were that the intervention was too time-consuming or too demanding. There were no significant differences between participants that adhered to the ACT-M and ACT-E intervention in the average number of weeks completed, number of emails sent, time spent weekly (in hours) on the self-help programme or the participant's satisfaction. They completed on average 7.1 weeks ($n=188$, range 4–8 weeks) and sent on average 7.4 emails ($n=200$, range 1–9 emails). The participants spent on average 4 h per week on the self-help programme. On the CSQ-8 participants evaluated the intervention with 3.2 (scale from 1 to 4) (s.d. = 0.44, $n=188$). On a scale from 1 to 10 the intervention was evaluated with a 7.7 (s.d. = 1.01, $n=194$).

Outcome of the interventions

Means and standard deviations for all the outcome measures as well as the results from the repeated-measures ANOVA are presented in Table 2. For all the outcome measures significant interactions were found [$F(2, 373)=18.96-38.92$, all $p<0.001$]. *Post-hoc* testing with the Tukey HSD test revealed that both the ACT-M and ACT-E participants were significantly improved from baseline to post-intervention compared with the W-L condition. The two intervention conditions did not significantly differ on improvement from baseline to post-intervention on any of the outcome measures. In Table 3 the results of the effect sizes between the conditions at post-intervention are presented. On all the outcome measures, moderate to high effect sizes were found for the ACT-M and ACT-E groups compared with the W-L condition.

Change from post-intervention to follow-up

Both experimental groups showed maintenance of the effects on the primary and secondary measures from post-intervention to follow-up. Significant reductions in EA were found from post-intervention to follow-up in the ACT-E group [$t(124)=-2.64$, $p=0.009$] and in the ACT-M group [$t(124)=-4.14$, $p<0.000$]. With regard to mindfulness, improvements were observed for the FFMQ-observe in the ACT-E group [$t(124)=-2.16$, $p=0.033$] and deteriorations in the FFMQ-describe in the ACT-M group [$t(124)=2.27$, $p=0.025$]. Overall these findings suggest that on all measures the effects were maintained or improved at the 3-month follow-up, except for the mindfulness facet 'describe'. Repeated-measures ANOVA revealed no significant differences between ACT-E and ACT-M on the change in outcome measures from post-intervention to follow-up.

Moderator analyses

A significant interaction was found for the condition by the baseline level of depressive symptoms ($<23=0$ and $\geq 23=1$) on the primary outcome ($\beta=-0.175$, $p=0.007$). Participants with a higher level of depressive symptoms at baseline showed greater reduction in depression after completion of the intervention compared with the W-L group. Between-group effect size at post-intervention was higher for people with higher levels of depression at baseline (Cohen's $d_{CES-D<23}=0.71$, Cohen's $d_{CES-D\geq 23}=1.08$). No significant interactions were found for the condition by gender, level of education (high or low), the level of anxiety symptoms at baseline ($<10=0$ and $\geq 10=1$) and the level of EA at baseline ($<40=0$ and $\geq 40=1$).

Clinically relevant change

The proportion of participants that reached a clinically significant change on the CES-D in the ACT-E group was 42/125 (34%) versus 7/126 (6%) in the W-L group [OR 8.60, 95% confidence interval (CI) 3.69–20.08, $p<0.001$, NNT=3.57]. In the ACT-M group, 49/125 (39%) reached a clinically significant change, which also resulted in a significant difference compared with the W-L group (OR 10.96, 95% CI 4.72–25.46, $p<0.001$, NNT=2.98). There were no significant differences between the experimental conditions on the proportions of clinically relevant change.

Medication use

Of the participants, 10 in the intervention groups (ACT-E=4, ACT-M=6) and three in the control group reported at baseline to have started more than 3 months previously with a psychopharmacological treatment (and were thus not excluded from the study). During the intervention seven participants started with medication (ACT-E=2, ACT-M=2, W-L=3) and seven participants stopped (ACT-E=1, ACT-M=5, W-L=1). Due to these small numbers and the equal distribution over the conditions, it is unlikely that the present findings can be explained by changes in medication use preceding or during the trial.

Discussion

Effects on psychological distress and positive mental health

Participants who received the guided ACT self-help intervention had significantly more reduction in depressive symptoms directly after the intervention compared with participants in the W-L group, and this reduction was sustained at the 3-month follow-up.

Table 2. Scores for all outcome measures and the results from the repeated-measures ANOVA

Outcome	Group	Score			ANOVA: F		
		Pre	Post	Follow-up	Time	Group	Time × group
Primary outcome							
CES-D	ACT-E	23.14 (6.48)	13.84 (7.55)	14.04 (8.62)	311.09***	11.50***	30.62***
	ACT-M	22.43 (6.65)	12.82 (6.99)	13.38 (7.78)			
	W-L	22.45 (6.68)	19.76 (8.48)				
Secondary outcomes							
HADS-A	ACT-E	9.36 (2.62)	6.22 (2.98)	5.92 (3.22)	273.00***	10.60***	38.53***
	ACT-M	9.67 (2.53)	6.02 (2.96)	5.91 (3.03)			
	W-L	9.33 (2.34)	8.69 (3.19)				
CIS	ACT-E	36.90 (11.05)	28.58 (11.40)	28.84 (12.77)	136.44***	8.06***	24.45***
	ACT-M	39.16 (9.80)	31.20 (11.85)	31.12 (12.03)			
	W-L	38.24 (10.58)	37.36 (11.51)				
MHC-SF-EM	ACT-E	3.27 (0.92)	4.10 (0.87)	4.17 (0.84)	193.27***	6.10**	30.04**
	ACT-M	3.42 (0.89)	4.26 (0.82)	4.21 (0.92)			
	W-L	3.43 (0.90)	3.56 (0.94)				
MHC-SF-SOC	ACT-E	2.79 (0.84)	3.48 (0.96)	3.51 (0.86)	151.52***	5.15**	19.77***
	ACT-M	2.81 (0.88)	3.46 (0.97)	3.51 (1.00)			
	W-L	2.78 (0.81)	2.92 (0.97)				
MHC-SF-PSY	ACT-E	3.20 (0.88)	4.04 (0.89)	4.01 (0.92)	207.03***	1.78	32.91***
	ACT-M	3.27 (0.94)	4.12 (0.93)	4.09 (0.94)			
	W-L	3.44 (0.77)	3.56 (0.99)				
Process measures							
FFMQ – observing	ACT-E	25.01 (5.24)	27.53 (4.72)	28.17 (4.29)	59.15***	4.85**	24.92***
	ACT-M	25.25 (5.11)	27.76 (4.56)	27.64 (4.92)			
	W-L	25.00 (5.16)	24.55 (5.99)				
FFMQ – describing	ACT-E	25.74 (6.42)	28.77 (5.41)	29.02 (5.21)	106.53***	4.57**	18.96***
	ACT-M	25.97 (6.00)	29.11 (5.72)	28.52 (5.82)			
	W-L	25.37 (6.22)	25.71 (6.37)				
FFMQ – acting with awareness	ACT-E	21.40 (4.84)	25.30 (4.67)	25.70 (5.25)	136.90***	8.20***	28.57***
	ACT-M	20.31 (5.24)	24.89 (5.08)	25.15 (5.36)			
	W-L	21.02 (4.80)	21.30 (5.23)				
FFMQ – non-judging	ACT-E	23.17 (5.27)	29.77 (5.47)	29.85 (5.96)	292.40***	11.71***	38.92***
	ACT-M	22.76 (5.49)	29.40 (5.77)	29.58 (5.83)			
	W-L	23.02 (5.38)	24.34 (6.48)				
FFMQ – non-reactivity	ACT-E	19.39 (3.55)	23.36 (3.66)	23.64 (3.90)	230.59***	14.34***	38.74***
	ACT-M	19.03 (3.88)	23.91 (3.70)	23.71 (4.01)			
	W-L	19.17 (3.89)	19.80 (4.49)				
AAQ-II	ACT-E	40.08 (7.93)	48.95 (8.55)	50.32 (9.75)	234.13***	6.51**	26.29***
	ACT-M	41.34 (8.96)	49.64 (9.59)	51.42 (9.71)			
	W-L	40.86 (8.83)	43.00 (10.27)				

ANOVA, Analysis of variance; CES-D, Center for Epidemiologic Studies Depression Scale; ACT-E, acceptance and commitment therapy with extensive email support; ACT-M, acceptance and commitment therapy with minimal email support; W-L, waiting list control; HADS-A, Hospital Anxiety and Depression Scale – anxiety subscale; CIS, Checklist Individual Strength; MHC-SF, Mental Health Continuum – short form; EM, emotional; SOC, social; PSY, psychological; FFMQ, Five Facet Mindfulness Questionnaire; AAQ-II, Acceptance and Action Questionnaire-II.

Data are given as mean (standard deviation).

** $p < 0.01$, *** $p < 0.001$.

The large effect sizes (Cohen's $d = 0.74$ – 0.89) found in this study are largely consistent with meta-analyses with guided self-help treatment for depression

(Gregory *et al.* 2004; Gellathly *et al.* 2007). As such, the results confirm earlier studies that showed that interventions for people with subthreshold depression are

Table 3. Effect sizes (Cohen's *d*) between conditions at post-intervention

	ACT-E v. W-L	ACT-M v. W-L
Primary outcome		
CES-D	0.74	0.89
Secondary outcomes		
HADS-A	0.80	0.87
CIS	0.77	0.53
MHC-SF-EM	0.60	0.79
MHC-SF-SOC	0.62	0.56
MHC-SF-PSY	0.51	0.58
Process measures		
FFMQ – observing	0.55	0.60
FFMQ – describing	0.52	0.56
FFMQ – acting with awareness	0.81	0.70
FFMQ – non-judging	0.91	0.82
FFMQ – non-reactivity	0.87	1.00
AAQ-II	0.63	0.70

ACT-E, Acceptance and commitment therapy with extensive email support; W-L, waiting list control; ACT-M, acceptance and commitment therapy with minimal email support; CES-D, Center for Epidemiologic Studies Depression Scale; HADS-A, Hospital Anxiety and Depression Scale – anxiety subscale; CIS, Checklist Individual Strength; MHC-SF, Mental Health Continuum – short form; EM, emotional; SOC, social; PSY, psychological; FFMQ, Five Facet Mindfulness Questionnaire; AAQ-II, Acceptance and Action Questionnaire-II.

effective in reducing depressive symptomatology (Cuijpers *et al.* 2007). Our study further showed that the intervention resulted in a particularly large effect size among participants with relatively high baseline levels of depressive symptomatology (Cohen's $d=1.08$). As the presence of clinically relevant depressive symptoms is known to be an important risk factor for clinical depression (Cuijpers & Smit, 2004), this outcome suggests that guided ACT self-help early intervention decreases the risk of developing a (future) full clinical depression. Moreover, the ACT intervention resulted in significant large reductions in anxiety and fatigue after the intervention and at the 3-month follow-up. This finding corroborates earlier studies that showed that ACT is effective in reducing anxiety and fatigue (Forman *et al.* 2007; Lappalainen *et al.* 2007; Bohlmeijer *et al.* 2011a).

The ACT intervention also significantly improved positive mental health, confirming previous research showing that ACT interventions promote emotional well-being (e.g. life satisfaction or quality of life) in patients with anxiety or depression (e.g. Forman *et al.* 2007; Lappalainen *et al.* 2007). Also psychological and

social well-being were improved, suggesting that an intervention that is aimed at increasing acceptance and value-based actions is supportive for a positive, engaged and meaningful life.

Effects on EA and mindfulness

Large effect sizes were found on EA at post-intervention and at follow-up further reductions were found in the experimental conditions. This finding corroborates previous studies that showed that EA can be substantially reduced by ACT. It indicates that the participants are more inclined to accept their undesirable feelings, thoughts or body sensations without trying to avoid the form or frequency (Hayes *et al.* 2006). This is also confirmed by the large effect sizes that were found on the FFMQ subscales 'non-judging' and 'non-reactivity'. This means that the participants have a more non-evaluative stance toward thoughts and feelings and allow them to come and go (Baer *et al.* 2006). On the other subscales of the FFMQ also large effects sizes were found in the experimental conditions after the intervention. Future research could examine the possibly mediating role of EA or mindfulness.

Level of support

No significant differences in the level of email support provided by counsellors were detected. This appears to be in line with earlier studies that have found that minimal contact with a counsellor is enough for a self-help programme to be as effective as face-to-face therapy (Cuijpers *et al.* 2010). Although our study was not powered to detect smaller effect size differences between two effective interventions (Cohen's $d < 0.40$) and thus more subtle differences could have been missed, the similarities in the effects were highly consistent. In both the minimal and extensive email support, moderate to large effect sizes were found on all outcome measures and the participants in both conditions were highly satisfied with the intervention. This implies that for this intervention minimal email support suffices and more extensive email counselling does not further improve outcomes. It should be noted that the intensity of email contact differed, not the frequency, as this was held constant in both conditions. Further research could examine whether even less frequent contact could be effective, such as merely providing a clear scheduled deadline for finishing the programme (e.g. Nordin *et al.* 2010) or as an online intervention with automatic feedback.

Limitations

This study has several limitations. The use of a W-L control group is suboptimal, because of a lack of

control for non-specific factors. Also, it would have strengthened the design if a longer follow-up period and a follow-up assessment for the control group had been included. With the present study it cannot be demonstrated that the effects at the 3-month follow-up were attributable to the ACT intervention and not, for example, to spontaneous remission.

Our participants were mainly highly educated females, so generalization of the results has to be made with prudence. However, reaching this group is not uncommon for self-help (e-health) interventions (e.g. Carlbring *et al.* 2007). Furthermore, moderator analyses showed that there were no significant intervention effects for gender or education, indicating that the intervention is potentially broadly applicable.

Another limitation is that only partial and limited use was made of a diagnostic instrument. The MINI was used for diagnosing severe depressive episodes with people who scored positively on the WSQ but who did not meet the exclusion criterion of a severe depressive disorder. Hence the sample in this study consisted of both people with clinically relevant depressive symptomatology and people with mild to moderate depressive disorders. However, this limitation can also be seen as a strength. Apparently, the inclusion of a heterogeneous population with regard to their range of depressive symptomatology did not negatively affect treatment effectiveness. This underscores the generalizability of the findings.

Implications and future directions

This study is the first to show that ACT is effective as an early intervention offered as a guided self-help programme to participants with mild to moderate depressive symptomatology. Offering an early intervention in a positive frame and as a self-help programme might be less stigmatizing for participants than the traditional mental health services. Our study has shown that the ACT intervention not only affected depressive symptomatology but also symptoms of anxiety and fatigue. Apparently, by targeting a general risk factor (EA), a broad spectrum of psychological distress can be reduced. This implies that a generic preventive intervention could extend to other mental disorders as well.

In our study positive mental health was a secondary outcome. There is a growing body of knowledge that underscores positive mental health as a major public health goal. Recent studies have shown that having positive mental health protects against the risk of mental illnesses (Keyes *et al.* 2010; Wood & Joseph, 2010). Furthermore, Fava *et al.* (2001) found that even when clinical symptoms are absent, low positive mental health can be seen as a risk factor for future

relapse among recovering patients. There is also growing evidence of interventions aiming at enhancing positive mental health both in clinical practice (e.g. Fava *et al.* 2005) and in public mental health (e.g. Fledderus *et al.* 2010). In future studies positive mental health could therefore be considered as a primary outcome of interventions aiming at people with low positive mental health.

Future research could examine whether a guided self-help ACT intervention is effective in reducing the onset of a major depressive disorder. Though a drop-out rate of 18% is acceptable and similar to other self-help studies (e.g. Seekles *et al.* 2011), it is an important issue to prevent drop-out from self-help interventions. Future studies could examine how lower drop-out rates can be achieved, for example, to tailor a self-help treatment according to the patient profile (e.g. Andersson *et al.* 2011). Furthermore, it could be examined whether the intervention could be tailored to people with a low social economic status, as this status is clearly associated with depression (Lorant *et al.* 2007). One important modification could be the use of comics to illustrate the main concepts and metaphors of the intervention.

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Declaration of Interest

None.

References

- Andersson G, Estling F, Jakobsson E, Cuijpers P, Carlbring P (2011). Can the patient decide which modules to endorse? An open trial of tailored Internet treatment of anxiety disorders. *Cognitive Behaviour Therapy* **40**, 57–64.
- Andrews G, Issakidis C, Sanderson K, Corry L, Lapsley H (2004). Utilizing survey data to inform public policy, comparison of the cost-effectiveness of treatment of ten mental disorders. *British Journal of Psychiatry* **84**, 526–533.
- Attkisson CC, Zwick R (1982). The client satisfaction questionnaire. Psychometric properties and correlations with service utilization and psychotherapy outcome. *Evaluation and Program Planning* **5**, 233–237.
- Baer RA, Smith GT, Hopkins J, Krietemeyer J, Toney L (2006). Using self-report assessment methods to explore facts of mindfulness. *Assessment* **13**, 27–45.

- Beekman AT, Deeg DJ, Van Limbeek J, Braam AW, De Vries MZ, Van Tilburg W (1997). Criterion validity of the Center for Epidemiologic Studies Depression scale (CESD), results from a community-based sample of older subjects in the Netherlands. *Psychological Medicine* 27, 231–235.
- Biglan A, Hayes SC, Pistorello J (2008). Acceptance and commitment, implications for prevention science. *Prevention Science* 9, 139–152.
- Bohlmeijer ET, Fledderus M, Rokx TAJJ, Pieterse ME (2011a). Efficacy of an early intervention based on acceptance and commitment therapy for adults with depressive symptomatology: evaluation in a randomized controlled trial. *Behaviour Research and Therapy* 9, 62–67.
- Bohlmeijer ET, Hulsbergen M (2008). *Voluit leven. Mindfulness of de kunst van het aanvoarden, nu als praktisch hulpboek [Living to the Full. Mindfulness or the Art of Acceptance, now as a Practical Help Book]*. Boom: Amsterdam.
- Bohlmeijer ET, Ten Klooster PM, Fledderus M, Veehof MM, Baer R (2011b). Psychometric properties of the Five Facet Mindfulness Questionnaire in depressed adults and development of a short form. *Assessment*. Published online 17 May 2011. doi:10.1177/1073191111408231.
- Bond FW, Hayes SC, Baer RA, Carpenter KC, Guenole N, Orcutt HK, Waltz T, Zettle RD (2011). Preliminary psychometric properties of the Acceptance and Action Questionnaire – II: a revised measure of psychological flexibility and acceptance. *Behavior Therapy*. doi:10.1016/j.beth.2011.03.00.
- Bouma J, Ranchor AV, Sanderman R, Van Sonderen E (1995). *Het meten van symptomen van depressie met de CES-D, een handleiding [Measuring Symptoms of Depression with the CES-D, A Guide]*. Noordelijk Centrum voor Gezondheidsvraagstukken: Groningen.
- Carlbring P, Gunnarsdottir M, Hedensjö L, Andersson G, Ekselius L, Furmark T (2007). Treatment of social phobia, randomized trial of Internet-delivered cognitive-behavioural therapy with telephone support. *British Journal of Psychiatry* 90, 123–128.
- Cook RJ, Sackett DL (1995). The number needed to treat, a clinically useful measure of treatment effect. *British Medical Journal* 310, 452–454.
- Cuijpers P, Donker T, Van Straten A, Li J, Andersson G (2010). Is guided self-help as effective as face-to-face psychotherapy for depression and anxiety disorders? A systematic review and meta-analysis of comparative outcome studies. *Psychological Medicine* 40, 1943–1957.
- Cuijpers P, Smit F (2004). Subthreshold depression as a risk indicator for major depressive disorder, a systematic review of prospective studies. *Acta Psychiatrica Scandinavica* 109, 325–331.
- Cuijpers P, Smit F, Van Straten A (2007). Psychological treatments of subthreshold depression, a meta-analytic review. *Acta Psychiatrica Scandinavica* 115, 434–441.
- Cuijpers P, Van Straten A, Smit F, Mihalopoulos C, Beekman A (2008). Preventing the onset of depressive disorders, a meta-analytic review of psychological interventions. *American Journal of Psychiatry* 65, 1272–1280.
- Cuijpers P, Van Straten A, Warmerdam L, Van Rooy MJ (2010). Recruiting participants for interventions to prevent the onset of depressive disorders. Possible ways to increase participation rates. *BMC Health Services Research* 10, 181.
- De Brey HA (1983). Cross-national validation of the client satisfaction questionnaire, the Dutch experience. *Evaluation and Program Planning* 6, 395–400.
- Dempster AP, Laird NM, Rubin DB (1977). Maximum likelihood from incomplete data via the EM algorithm. *Journal of the Royal Statistical Society Series B Methodological* 39, 1–38.
- Donker T, Van Straten A, Marks I, Cuijpers P (2009). A brief web-based screening questionnaire for common mental disorders, development and validation. *Journal of Medical Internet Research* 11, 19.
- Fava GA, Rafanelli C, Ottolini F, Ruini C, Cazzaro M, Grandi S (2001). Psychological well-being and residual symptoms in remitted patients with panic disorder and agoraphobia. *Professional Psychology, Research and Practice* 31, 899–905.
- Fava GA, Ruini C, Rafanelli C, Finos L, Salmaso L, Mangelli L, Sirigatti S (2005). Well-being therapy of generalized anxiety disorder. *Psychotherapy Psychosomatic* 74, 26–30.
- Fledderus M, Bohlmeijer ET, Smit F, Westerhof GJ (2010). Mental health promotion as a new goal in public mental health care. A randomized controlled trial of an intervention enhancing psychological flexibility. *American Journal of Public Health* 100, 2372–2378.
- Forman EM, Herbert JD, Moitra E, Yeomans PD, Geller PA (2007). Randomized controlled effectiveness trial of acceptance and commitment therapy and cognitive therapy for anxiety and depression. *Behavior Modification* 31, 772–799.
- Gellathly J, Bower P, Hennessy S, Richards D, Gilbody S, Lovell K (2007). What makes self-help interventions effective in the management of depressive symptoms? Meta-analysis and meta-regression. *Psychological Medicine* 37, 1217–1228.
- Gregory R, Canning S, Lee T, Wise J (2004). Cognitive bibliotherapy for depression, a meta-analysis. *Professional Psychology, Research and Practice* 35, 275–280.
- Haringsma R, Engels GI, Beekman ATF, Spinhoven P (2004). The criterion validity of the Center for Epidemiological Studies Depression Scale (CES-D) in a sample of self-referred elders with depressive symptomatology. *International Journal of Geriatric Psychiatry* 19, 558–563.
- Hayes SC, Luoma JB, Bond FW, Masuda A, Lillis J (2006). Acceptance and commitment therapy. Model, processes and outcomes. *Behaviour Research and Therapy* 44, 1–25.
- Jacobs N, Kleen M, De Groot F, A-Tjak J (2008). Het meten van experiëntiële vermijding. De Nederlandstalige versie van de Acceptance and Action Questionnaire-II (AAQ-II) [Measuring experiential avoidance. Dutch translation of the Acceptance and Action questionnaire-II (AAQ-II)]. *Gedragstherapie* 41, 349–361.
- Jacobson NS, Truax P (1991). Clinical significance, a statistical approach to defining meaningful change in

- psychotherapy research. *Journal of Consulting and Clinical Psychology* **59**, 12–19.
- Kabat-Zinn J** (1990). *Full Catastrophe Living, Using the Wisdom of Your Body and Mind to Face Stress, Pain, and Illness*. Delacorte: New York.
- Kabat-Zinn J** (1994). *Wherever You Go, There You Are, Mindfulness Meditation in Everyday Life*. Hyperion: New York.
- Keyes CLM** (2005). Mental illness and/or mental health? Investigating axioms of the complete state model of health. *Journal of Consulting and Clinical Psychology* **73**, 539–548.
- Keyes CLM, Dhingra SS, Simoes EJ** (2010). Change in level of positive mental health as a predictor of future risk of mental illness. *American Journal of Public Health* **100**, 2366–2371.
- Lamers SMA, Westerhof GJ, Bohlmeijer ET, Ten Klooster PM, Keyes CLM** (2010). Evaluating the psychometric properties of the Mental Health Continuum-Short Form (MHC-SF). *Journal of Clinical Psychology* **67**, 99–110.
- Lappalainen R, Lehtonen T, Skarp E, Taubert E, Ojanen M, Hayes SC** (2007). Preliminary controlled effectiveness trial: the impact of CBT and ACT models using psychology trainee. *Behavior Modification* **31**, 488–511.
- Lipsey MW, Wilson DB** (1993). The efficacy of psychological, educational and behavioral treatment. Confirmation from meta-analysis. *American Psychologist* **48**, 1181–1209.
- Lorant V, Croux C, Weich S, Deliège D, Mackenbach J, Anseau M** (2007). Depression and socio-economic risk factors: 7-year longitudinal population study. *British Journal of Psychiatry* **190**, 293–298.
- Nordin S, Carlbring P, Cuijpers P, Andersson G** (2010). Expanding the limits of bibliotherapy for panic disorder: randomized trial of self-help without support but with a clear deadline. *Behavior Therapy* **41**, 267–276.
- Olsson I, Mykletun A, Dahl AA** (2005). The Hospital Anxiety and Depression rating scale, a cross-sectional study of psychometrics and case findings abilities in general practice. *BMC Psychiatry* **5**, 46.
- Radloff LS** (1977). The CES-D scale, a self-report depression scale for research in the general population. *Applied Psychological Measurement* **1**, 385–401.
- Seekles W, van Straten A, Beekman A, van Marwijk H, Cuijpers P** (2011). Effectiveness of guided self-help for depression and anxiety disorders in primary care: a pragmatic randomized controlled trial. *Psychiatry Research* **187**, 113–120.
- Sheehan DV, Lecrubier Y, Sheehan KH, Amorim P, Janav J, Weiller E, Hergueta T, Baker R, Dunbar GC** (1998). The Mini-International Neuropsychiatric Interview (MINI). The development and validation of a structured diagnostic psychiatric interview for DSM-IV and ICD-10. *Journal of Clinical Psychiatry* **59**, 22–33.
- Smit F, Ederveen A, Cuijpers P, Deeg D, Beekman A** (2006). Opportunities for cost-effective prevention of late-life depression. An epidemiological approach. *Archives of General Psychiatry* **63**, 290–296.
- Spinhoven PH, Ormel J, Sloekers PPA, Kempen GIJM, Speckens AEM, Van Hemert AM** (1997). A validation study of the Hospital Anxiety and Depression Scale (HADS) in different groups of Dutch subjects. *Psychological Medicine* **27**, 363–370.
- Vercoulen JHM.M, Alberts M, Bleijenberg G** (1999). The Checklist Individual Strength (CIS). *Gedragstherapie* **32**, 131–136.
- Wood AM, Joseph S** (2010). The absence of positive psychological (eudemonic) well-being as a risk factor for depression, a ten year cohort study. *Journal of Affective Disorders* **122**, 213–217.
- Zigmond AS, Snaith RP** (1983). The Hospital Anxiety and Depression Scale. *Acta Psychiatrica Scandinavica* **67**, 361–370.